

Econometrics

Lecture 4

Nathaniel Higgins

ERS and JHU

26 September 2011

- Any questions?

Regression

How to think about regression

- We might think that y is related to x by a certain function
- We are interested in how y is related to x *on average*
- This notion is captured in what is called the conditional expectation function

$$E(y|x) = \beta_0 + \beta_1 x$$

- The conditional expectation function gives the expected value of y (the variable that is being **affected**), *conditional on x*
- The term *conditional on x* means “given what we know about x ”

Regression

Conditional expectation

- The term *conditional on x* means “given what we know about x ”
- Forget about regressions for just a moment. Suppose we know that the expected value of income in the U.S. for an individual is \$30k
- If you pick a random person out of the population (totally random!) and you have to guess their income, what would your guess be? Probably \$30k
- Now suppose that before you guessed, you were given one extra fact about the person: they live in Malibu. Would you guess \$30k? Probably not. You would guess higher
- The first guess (before you knew anything about the randomly drawn person) is akin to $E(y)$
- The second guess (after you knew that they lived in Malibu) is akin to $E(y|x)$, where x represents where the individual lives

Regression

Conditional expectation

- Get it?
- The conditional expectation is what we expect given all the knowledge and information that we have
- This is a very useful concept
- Remember that this was the representation of the conditional expectation function that we had

$$E(y|x) = \beta_0 + \beta_1 x$$

- If we had all the information — if nothing was left out of the equation explaining y — then this would be enough
- For every value of x it would be straightforward to know what y was. If $x = 5$, then $y = \beta_0 + \beta_1 * 5$. Period.
- But this is unrealistic. We don't know everything about how y is determined. So there is some extra randomness, which we add into the model with u

$$E(y|x) = \beta_0 + \beta_1 x + u$$



- We start with a variable that is “caused” or “determined.” We call it y .
- We think that y is determined in part by some variable x .
- We create some model of how y is determined by a function of x

$$y = f(x)$$

- We typically estimate $f(x)$ by a linear function (because they are simple and work pretty darn well)
 - The typical example is $f(x) = \beta_0 + \beta_1 x$
- We get some data and estimate the relationship so that we get a prediction of what y is for every value of x

$$\hat{y} = \widehat{f(x)} = \hat{\beta}_0 + \hat{\beta}_1 x$$

- How do we get these predictions?

Regression

Summary

- We get these predictions by selecting our favorite values for $\widehat{\beta}_0$ and $\widehat{\beta}_1$. How do we select our favorite values?
- Well, we look at our data and it looks like this

y	x
y_1	x_1
y_2	x_2
y_3	x_3
\vdots	\vdots
y_n	x_n

Regression

Summary

y	x
y_1	x_1
y_2	x_2
y_3	x_3
\vdots	\vdots
y_n	x_n

Regression

Summary

y	x	
y_1	x_1	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$
y_2	x_2	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$
y_3	x_3	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$
\vdots	\vdots	
y_n	x_n	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$

Regression

Summary

y	x		
y_1	x_1	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$	\widehat{y}_1
y_2	x_2	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$	\widehat{y}_2
y_3	x_3	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$	\widehat{y}_3
\vdots	\vdots		\vdots
y_n	x_n	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$	\widehat{y}_n

Regression

Summary

y	x			“errors”	
y_1	x_1	\rightarrow	$(\widehat{\beta}_0, \widehat{\beta}_1)$	\rightarrow	\widehat{y}_1
y_2	x_2	\rightarrow	$(\widehat{\beta}_0, \widehat{\beta}_1)$	\rightarrow	\widehat{y}_2
y_3	x_3	\rightarrow	$(\widehat{\beta}_0, \widehat{\beta}_1)$	\rightarrow	\widehat{y}_3
\vdots	\vdots				\vdots
y_n	x_n	\rightarrow	$(\widehat{\beta}_0, \widehat{\beta}_1)$	\rightarrow	\widehat{y}_n

Regression

Summary

y	x				"errors"	
y_1	x_1	\rightarrow	$(\widehat{\beta}_0, \widehat{\beta}_1)$	\rightarrow	\widehat{y}_1	$(y_1 - \widehat{y}_1)^2$
y_2	x_2	\rightarrow	$(\widehat{\beta}_0, \widehat{\beta}_1)$	\rightarrow	\widehat{y}_2	$(y_2 - \widehat{y}_2)^2$
y_3	x_3	\rightarrow	$(\widehat{\beta}_0, \widehat{\beta}_1)$	\rightarrow	\widehat{y}_3	$(y_3 - \widehat{y}_3)^2$
\vdots	\vdots				\vdots	
y_n	x_n	\rightarrow	$(\widehat{\beta}_0, \widehat{\beta}_1)$	\rightarrow	\widehat{y}_n	$(y_n - \widehat{y}_n)^2$

Regression

Summary

y	x			“errors”
y_1	x_1	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$	\widehat{y}_1	$(y_1 - \widehat{y}_1)^2$
y_2	x_2	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$	\widehat{y}_2	$(y_2 - \widehat{y}_2)^2$
y_3	x_3	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$	\widehat{y}_3	$(y_3 - \widehat{y}_3)^2$
\vdots	\vdots		\vdots	
y_n	x_n	$\rightarrow (\widehat{\beta}_0, \widehat{\beta}_1) \rightarrow$	\widehat{y}_n	$(y_n - \widehat{y}_n)^2$

$$\sum_{i=1}^n (y_i - \widehat{y}_i)^2$$

- We select the values of $\widehat{\beta}_0$ and $\widehat{\beta}_1$ that minimize the squared difference between y and \widehat{y}

$$(y - \widehat{y})^2$$

- That's it!
- Now we have two things:
 - 1 Point estimates of the true parameters β_0 and β_1
 - 2 Predictions of y for every value of x in our data
- Before we explore anything else, let's try some basic regressions

- Load up some data from Wooldridge

Stata code

```
use ceosall.dta
```

```
* What is in the data?
```

```
* What are the basic statistics of the data?
```

- Load up some data from Wooldridge

Stata code

```
use ceosall.dta
* What is in the data?
describe
* What are the basic statistics of the data?
```


- Load up some data from Wooldridge

Stata code

```
use ceosall.dta
* What is in the data?
describe
* What are the basic statistics of the data?
sum
```

Stata code

- * We are interested in the relationship
- * between CEO salary (salary) and return on
- * investment (roe)
- * Let's look at how they appear to be related
- * by examining a simple scatter of the data
- scatter salary roe
- * What is the first thing that you notice?

Stata code

```
* Eliminate some of the outliers  
* so that we can get a better look  
scatter salary roe if (salary < 5000)
```

Stata code

```
* Now that we have a rough idea of what  
* the data look like, let's run a model  
reg salary roe
```

Stata code

```
* Let's draw the model on top of the data
* First, let's get the predictions of the
* model
predict yhat
* Now plot yhat against roe
```

Stata code

```
* Let's draw the model on top of the data
* First, let's get the predictions of the
* model
predict yhat
* Now plot yhat against roe
scatter salary roe if salary < 5000 ///
|| scatter yhat roe
```

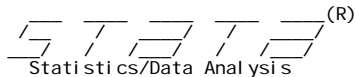
Stata code

```
* One last thing. We might want to see a line  
* (rather than a scatter) of the predicted  
* values  
scatter salary roe if salary < 5000 ///  
|| line yhat roe
```

Summing up

- What is the interpretation of our estimates $\widehat{\beta}_0$ and $\widehat{\beta}_1$?
- Now, that's all well and good. But you might have noticed there are some other things printed to the screen when you type `reg salary roe`.
- We move next to these other things.
- Look again at the results on your screen

The results on your screen



```
. reg salary roe
```

Source	SS	df	MS
Model	5166419.04	1	5166419.04
Residual	386566563	207	1867471.32
Total	391732982	208	1883331.64

Number of obs = **209**
F(1, 207) = **2.77**
Prob > F = **0.0978**
R-squared = **0.0132**
Adj R-squared = **0.0084**
Root MSE = **1366.6**

salary	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
roe	18.50119	11.12325	1.66	0.098	-3.428196 40.43057
_cons	963.1913	213.2403	4.52	0.000	542.7902 1383.592

The results on your screen

- There is a lot there. Some of the stuff is ignorable for now. But let's start with a few things you can immediately understand
- Look in the upper-right. `Number of obs` is n : the number of observations
- We also see the coefficients corresponding to `roe` ($\widehat{\beta}_1$) and `_cons` ($\widehat{\beta}_0$)
- These are the things we have covered so far. The next thing we want to look at are the numbers under the column heading `SS` in the upper-left of the results

Sum of squares

TSS

- SS stands for sum of squares
- There are 3 different “SS” entries:
 - 1 Model
 - 2 Residual
 - 3 Total
- Let's deal with the Total SS first
 - The Total SS is the sum of squared differences between y and its mean \bar{y}
 - This is called the Total SS (or TSS) because it represents all the variation of y
 - This variation in y is what the model is trying to explain!
 - When we run our econometric model, we are trying to explain movements in y with movements in $x \rightarrow$ we cannot explain more than the total variation in y that we observe
- $TSS = \sum (y_i - \bar{y}_i)^2$

Sum of squares

TSS

- You can picture the variation in y as being represented by a circle — the bigger the circle, the more variation there is to explain
- FIGURE

Sum of squares

MSS/SSE

- Just to confuse you, Stata and Wooldridge don't agree on what to call the next thing
- Stata: Model SS (MSS); Wooldridge: Explained SS (SSE)
- Either way, they mean the same thing: the sum of the difference between \hat{y} and its mean $\bar{\hat{y}}$
- The variation in our predicted y 's is less (often much less) than the variation in the original y 's
- This is because our model doesn't predict all of the movement in y — our model isn't all powerful
- Picture the MSS (or SSE) as a circle also

Sum of squares

R-squared

- The overlap of these two circles represents the proportion of the total variation in y that is explained by our model (explained by x)
- There is a single number that summarizes this information:

$$R^2 = \frac{MSS}{TSS} \text{ or } \frac{SEE}{TSS}$$

- Try calculating for yourself

Stata

Wooldridge

Stata code

```
* Notice that lots of good stuff is saved  
* whenever we run a regression. Check out the  
* bottom of the help browser after typing:  
help reg
```

Stata code

```
* Notice that lots of good stuff is saved
* whenever we run a regression. Check out the
* bottom of the help browser after typing:
help reg
scalar TSS = e(mss) + e(rss)
scalar rSq = e(mss)/TSS
```


Sum of squares

R-squared

- You should notice that the number you got for R^2 is the same as what is printed in the upper right-hand column of your Stata results after you executed the command

```
reg salary roe
```

A word about functional form

- Everything we have learned so far has been about linear models

$$y = \beta_0 + \beta_1 x + u$$

- We say that the model above is *linear*
- What would a *nonlinear* model look like?

$$y = \left(\frac{\beta_1 x}{\beta_0 + x} \right)^u$$

- I just made that up. It doesn't mean anything, but it is *nonlinear*

A word about functional form

- We don't know what to do with a model like that yet
- There is a thing called nonlinear least squares (NLLS, as opposed to OLS). But we aren't there yet.
- There is a small set of nonlinear functional forms that you *can* work with
- ...because you can transform them into linear functional forms
- Consider an exponential function: check out the Stata example
- (you've heard of “exponential growth” right?)

Stata code

```
* Going to create some data with 100 obs
set obs 100
* Generate some random x data
gen x = 3*runiform()
* Generate random unobservable data
gen u = rnormal()
* Generate a nonlinear y
gen y = exp(x) + u
* Look what happens when you regress y on x
* when the relationship is nonlinear
scatter y x || lfit y x
```

Stata code

```
* How to solve the problem? Transform to  
* a linear model  
gen logY = log(y)  
* Now regress the new variable on x  
scatter logY x || lfit logY x
```

Why do we like OLS so much?

Properties of the OLS estimator

- It is linear (simplicity; works well in lots of situations)
- Unbiased
- As it happens, the OLS estimator is the *best* linear unbiased estimator that we could possibly come up with!
- There is an acronym for this: BLUE
- The linearity of OLS is inherent; the unbiasedness and the “bestness” of OLS are based on some assumptions that we should state and discuss

Why do we like OLS so much?

Assumptions

- Assumptions necessary to insure that OLS is BLUE
- Unobservables are conditionally independent from x

$$E(u|x) = 0$$

- y and x are randomly sampled (“i.i.d”)
- y and x have finite variance

Why do we like OLS so much?

Unbiasedness

- If we make these assumptions, we know that OLS is *unbiased*
- Recall what *unbiased* means:
 - An estimator is unbiased if it is equal to the true value, *on average*
 - Remember, our estimate of β_1 , which we call $\widehat{\beta}_1$ is itself a random variable, so it doesn't take on the same value every time
 - If β_1 is a single true number, and $\widehat{\beta}_1$, is a random variable, then sometimes $\widehat{\beta}_1$ takes on values different from β_1 . That is, sometimes $\widehat{\beta}_1$ is *wrong*
 - But if $\widehat{\beta}_1$ is right on average, then $\widehat{\beta}_1$ is unbiased

Why do we like OLS so much?

Unbiasedness

- The proof of unbiasedness is beyond the scope of this class
- If you plan to continue with econometrics at a high level, try to work through the proof on page 50 of Wooldridge
- A lot of the proof is “just math,” so to speak, but there is a wee bit of intuitive content

Why do we like OLS so much?

Unbiasedness

- The formula for the expectation of $\hat{\beta}$ works out to

$$E(\hat{\beta}) = \beta + f(x, u)$$

- The expected value of our estimator is equal to the true value of the parameter, plus some other stuff
- The other stuff is a function of the x data and the unobservables
- If the x data and the unobservables are not related at all (remember the assumption that $E(u|x) = 0$?) then the term $f(x, u)$ is equal to zero and we are left with

$$E(\hat{\beta}) = \beta$$

- We have several other properties of the OLS estimator to talk about once we get to the multivariate model
- The last thing we need to discuss before we take that leap is *causality*
- We have been looking at a simple model that looks at how much of the movement in y can be explained by x
- When we say that y is *explained by* x , it is natural to interpret this as causality: x does this, so y does that
- The interpretation of the OLS model is not, however, always causal
- The assumption that $E(u|x) = 0$ is key to this issue

Causality

An illustrative example

- Do hospitals make people feel better (make them less sick)?
- Said another way: does a hospital visit increase health?
- Seems like a straightforward question with an intuitively straightforward answer

Causality

An illustrative example

- Suppose we set up a survey; we have survey administrators set up at Dupont circle
- Survey administrators ask people how they are feeling; let's say they report how they are feeling on a 100 point scale (`health`)
- Survey administrators also ask people how many times they have visited the hospital in the last month (`hospital`)
- We run our favorite model (which is easy to select, since we only know 1 right now!)

$$\text{health} = \beta_0 + \beta_1 \text{hospital} + u$$

- We find that $\widehat{\beta}_1 < 0$!
- What is going on?

Causality

An illustrative example

- Why do people go to the hospital?

Causality

An illustrative example

- Why do people go to the hospital?
- Because they are sick!

Causality

An illustrative example

- Why do people go to the hospital?
- Because they are sick!
- So let's say that folks who go to the hospital (because they are sick) go to the hospital feeling like they about a 20 (on a 100 point scale)
- They visit the hospital and the hospital makes them feel better
- Say the hospital makes them feel 20 points better (not bad!)
- So what's the average health of a hospital-goer now? A: 40
- What do you think the average answer of a non-hospital-goer would be?

Causality

An illustrative example

- The folks who didn't go to the hospital were fine to begin with
- Suppose they had an average score of 60

Causality

An illustrative example

- The folks who didn't go to the hospital were fine to begin with
- Suppose they had an average score of 60
- If we run a regression with this data, we are comparing sick people to healthy people in order to try to find out if hospitals make people feel better
- So you see the problem intuitively
- *This* is the embodiment of a situation when $E(u|x) \neq 0$

Causality

An illustrative example

- The folks who didn't go to the hospital were fine to begin with
- Suppose they had an average score of 60
- If we run a regression with this data, we are comparing sick people to healthy people in order to try to find out if hospitals make people feel better
- So you see the problem intuitively
- *This* is the embodiment of a situation when $E(u|x) \neq 0$
- The unobservable term u contained a key piece of information that wasn't in our data: how the person was feeling when they went to the hospital in the first place, i.e. whether or not they were sick
- The unobservables were related to the x variable (number of hospital visits last month)

Causality

An illustrative example

- The folks who didn't go to the hospital were fine to begin with
- Suppose they had an average score of 60
- If we run a regression with this data, we are comparing sick people to healthy people in order to try to find out if hospitals make people feel better
- So you see the problem intuitively
- *This* is the embodiment of a situation when $E(u|x) \neq 0$
- The unobservable term u contained a key piece of information that wasn't in our data: how the person was feeling when they went to the hospital in the first place, i.e. whether or not they were sick
- The unobservables were related to the x variable (number of hospital visits last month)
- We can't have this if we want to give the econometric model a causal interpretation

- Experimental procedure
 - Suppose our population of interest is Washington DC ...

- Experimental procedure
 - Suppose our population of interest is Washington DC ...
 - ... so we drive around DC in a van with a sliding door, and we snatch people off the street

- Experimental procedure
 - Suppose our population of interest is Washington DC ...
 - ... so we drive around DC in a van with a sliding door, and we snatch people off the street
 - Then we take everyone to a warehouse, and we sort them into two categories, *randomly*: those who will go to the hospital, and those who will not

- Experimental procedure
 - Suppose our population of interest is Washington DC ...
 - ... so we drive around DC in a van with a sliding door, and we snatch people off the street
 - Then we take everyone to a warehouse, and we sort them into two categories, *randomly*: those who will go to the hospital, and those who will not
 - Then, after the "treatment" group has gone to the hospital and the "control" group has not, we take our measurements
 - We ask all the subjects to rank how good they feel, on a scale of 1 to 5

Wrapping-up

- There are a few topics covered in Ch 2 of Wooldridge that we will postpone until we cover multiple regression
- Variance of the estimators (standard error of the estimators)
- Homoskedasticity and its evil twin heteroskedasticity
- Regression through the origin
- We have covered everything else in Ch 2 and are ready for multiple regression

Next time

- Multiple regressors (multiple different x 's)
- Chapter 3 material
- Homework: Read 3.1, 3.2, 3.3